## Bayesian Limits: The CDF7117 Study

## Joint Physics Meeting, Feb 4 2005 Joel Heinrich

The plan for this talk is:

- Present a condensation of CDF7117, an investigation of the Bayesian approach to setting cross section upper limits.
- Mention available Bayesian limit calculating software.
- Concluding observations.

The problem: n events are observed in an experiment with Poisson probability

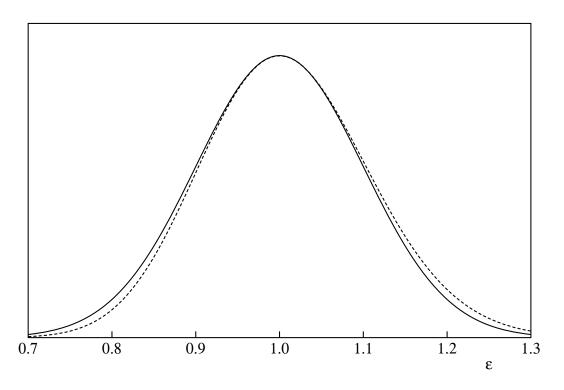
$$P(n) = \frac{e^{-(\epsilon s + b)}(\epsilon s + b)^n}{n!}$$

We wish to set an upper limit on the cross section s. The acceptance  $\epsilon$  and the expected background b are known with some uncertainty. For simplicity, CDF7117 deals primarily with the case  $\sigma_b = 0$ ; the extension to  $\sigma_b > 0$  is straightforward.

The Bayesian approach then requires priors for s and  $\epsilon$  (and for b, when  $\sigma_b > 0$ ).

#### The Acceptance Prior

The acceptance (here, the product of efficiency and luminosity) is normally measured in a separate, subsidiary, measurement or calculation. For CDF7117, we defined a specific subsidiary measurement precisely, which led to a gamma distribution prior for the acceptance. For  $\sigma_{\epsilon}=10\%$ , the gamma and Gaussian are quite similar:



acceptance prior p.d.f.

gamma: dashed

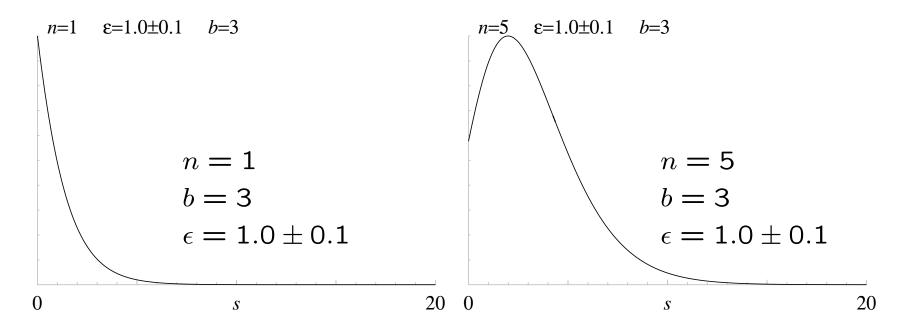
Gaussian: solid

#### The Cross Section Prior

We take the prior for s to be flat for  $s \ge 0$ . This is well behaved when combined with a gamma prior for  $\epsilon$ , but badly behaved with a truncated Gaussian  $\epsilon$  prior. Luc Demortier will discuss a promising solution to this problem that involves a different cross section prior (the usual "cutoff" solution is unpleasant).

#### Marginalized Posterior p.d.f. for s

From the likelihood and the priors, we obtain the joint posterior p.d.f. for the cross section and nuisance parameters. Integrating over the nuisance parameters then yields the marginalized posterior p.d.f. for the cross section.



The posterior is the "answer" in the Bayesian approach; one should look at its shape before proceeding. One obtains an upper limit at credibility level  $\beta$  by finding  $s_{\rm u}$  such that fraction  $\beta$  of the posterior is  $< s_{\rm u}$ .

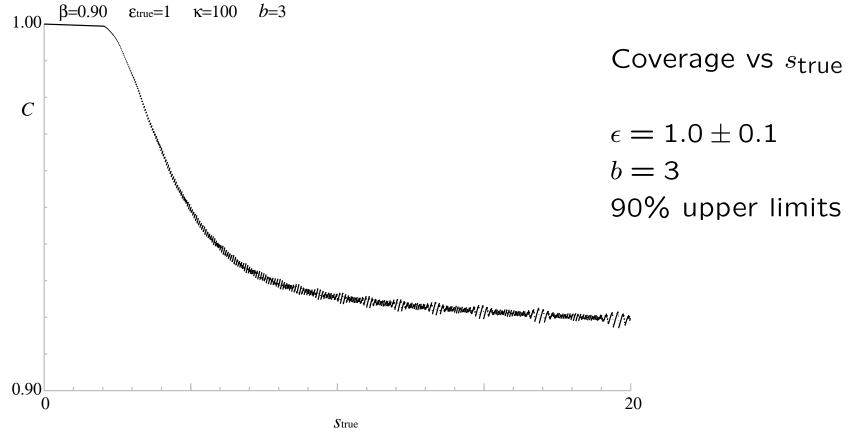
In the Bayesian approach, when n=0, the posterior for s, and hence the limit, is completely independent of the expected background b (and  $\sigma_b$ ). Conceptually, when n=0, we know that there were exactly zero signal events and exactly zero background events, yielding effectively perfect separation of signal from background in this special case. (The same conclusion does not necessarily hold in a frequentist approach.)

### 90% Upper Limits

|    | $\epsilon = 1.0 \pm 0.1$ |         |         |         |         |         |         |         |
|----|--------------------------|---------|---------|---------|---------|---------|---------|---------|
| n  | b = 0                    | 1       | 2       | 3       | 4       | 5       | 6       | 7       |
| 0  | 2.3531                   | 2.3531  | 2.3531  | 2.3531  | 2.3531  | 2.3531  | 2.3531  | 2.3531  |
| 1  | 3.9868                   | 3.3470  | 3.0620  | 2.9019  | 2.8000  | 2.7297  | 2.6783  | 2.6391  |
| 2  | 5.4669                   | 4.5520  | 3.9676  | 3.6026  | 3.3623  | 3.1953  | 3.0736  | 2.9816  |
| 3  | 6.8745                   | 5.8618  | 5.0463  | 4.4644  | 4.0571  | 3.7666  | 3.5534  | 3.3922  |
| 4  | 8.2380                   | 7.1964  | 6.2451  | 5.4751  | 4.8914  | 4.4569  | 4.1313  | 3.8832  |
| 5  | 9.5714                   | 8.5213  | 7.5063  | 6.6022  | 5.8579  | 5.2719  | 4.8180  | 4.4660  |
| 6  | 10.8826                  | 9.8288  | 8.7885  | 7.8047  | 6.9344  | 6.2066  | 5.6184  | 5.1499  |
| 7  | 12.1766                  | 11.1203 | 10.0703 | 9.0460  | 8.0904  | 7.2450  | 6.5289  | 5.9387  |
| 8  | 13.4570                  | 12.3984 | 11.3441 | 10.3014 | 9.2952  | 8.3635  | 7.5374  | 6.8300  |
| 9  | 14.7261                  | 13.6655 | 12.6085 | 11.5575 | 10.5247 | 9.5365  | 8.6250  | 7.8142  |
| 10 | 15.9858                  | 14.9233 | 13.8641 | 12.8090 | 11.7630 | 10.7415 | 9.7701  | 8.8758  |
| 11 | 17.2375                  | 16.1732 | 15.1121 | 14.0542 | 13.0017 | 11.9621 | 10.9525 | 9.9966  |
| 12 | 18.4823                  | 17.4163 | 16.3533 | 15.2934 | 14.2371 | 13.1881 | 12.1560 | 11.1582 |
| 13 | 19.7210                  | 18.6535 | 17.5887 | 16.5269 | 15.4682 | 14.4139 | 13.3692 | 12.3452 |
| 14 | 20.9545                  | 19.8854 | 18.8191 | 17.7554 | 16.6946 | 15.6373 | 14.5856 | 13.5459 |
| 15 | 22.1832                  | 21.1127 | 20.0448 | 18.9795 | 17.9169 | 16.8572 | 15.8014 | 14.7528 |
| 16 | 23.4078                  | 22.3359 | 21.2665 | 20.1996 | 19.1353 | 18.0737 | 17.0151 | 15.9612 |
| 17 | 24.6286                  | 23.5553 | 22.4845 | 21.4161 | 20.3502 | 19.2868 | 18.2261 | 17.1689 |
| 18 | 25.8459                  | 24.7714 | 23.6992 | 22.6294 | 21.5619 | 20.4969 | 19.4344 | 18.3747 |
| 19 | 27.0601                  | 25.9844 | 24.9109 | 23.8397 | 22.7708 | 21.7042 | 20.6400 | 19.5784 |
| 20 | 28.2715                  | 27.1946 | 26.1199 | 25.0474 | 23.9770 | 22.9090 | 21.8432 | 20.7799 |

#### Coverage

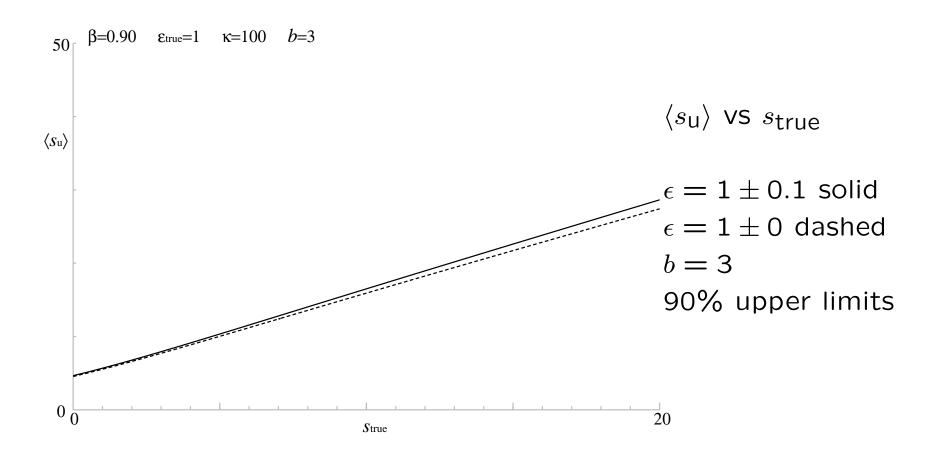
The coverage probability, a function of the true value of the cross section (and nuisance parameters), is the (frequentist) probability that a repetition of the experiment will yield a limit that includes the true value of the cross section.



Bayesian methods achieve average coverage, as defined in CDF7117. Frequentists often want minimum coverage for this Poisson case, in which constant coverage is not possible. With an (improper) flat cross section prior, it appears that both average coverage and minimum coverage are achieved for upper limits in the Bayesian approach.

(We use credibility level in a Bayesian context, and confidence level in a frequentist context.)

Another quantity of interest is the expected limit, or sensitivity, as a function of the true cross section.



# Currently Available Bayesian Limit Software from CDF www-cdf.fnal.gov/physics/statistics/statistics\_software.html

- C code from CDF7117 study: gamma acceptance and back-ground priors, various cross section priors. User can obtain posterior p.d.f., integral of posterior, upper limit at requested credibility level. Very fast execution time. User guide is CDF7232.
- bayes.f by John Conway. Traditional approach with truncated Gaussian priors for acceptance and background, flat prior with cutoff for cross section. Returns upper limit at requested credibility level. Slower because of MC integration. See CDF4476 by John Conway and Kaori Maeshima.

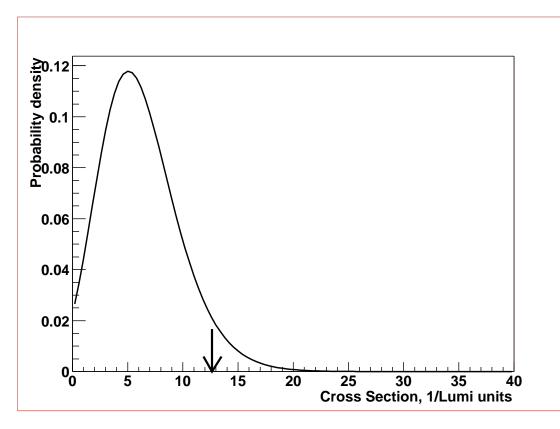
Other code available privately from John Conway:

bcorr.f - multi-channel counting experiment with correlated and uncorrelated uncertainties (between channels, and between signal and background) a la CDF 6428

fit.f - likelihood fit to spectrum with signal plus n background sources, including uncorrelated and correlated uncertainties on signal, background (implemented as nuisance parameters in the fit) described in CDF 6888

(fit.f is a profile likelihood approach.)

DØ (www-d0.fnal.gov/~hobbs/limit\_calc.html) has a web based menu driven product which is interesting—similar to bayes.f, but produces limit (95% only) and posterior plot online:



Data: 10

Background: 5 +- 1

Efficiency: 1.0 +- 0.1

Luminosity: 1.0 +- 0.0

The cross section 95% CL upper limit is 12.666

### **Concluding Observations**

- With a flat cross section prior and a gamma prior for acceptance, the Bayesian upper limit scheme exhibits no known defects. (Common defects are listed in CDF7117.)
- CDF7117 carries through all calculations analytically, which results in very fast computations. (Convenient for extensive coverage/sensitivity calculations.)
- Software is available for several combinations of priors.
- The gamma prior is more realistic than Gaussian for larger uncertainties (for small uncertainties they are very similar).
- Luc Demortier's correlated prior approach, described in the next talk, improves robustness, and represents a trivial modification to the existing code.